

# Mortgage Leverage and House Prices

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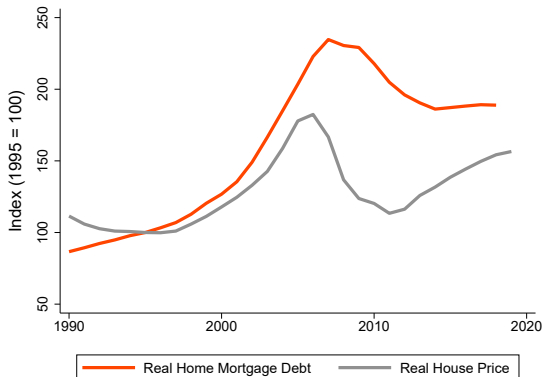
Rice University

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# How do house prices respond to lending standards?

- Lenders limit mortgage payment size relative to income.
- How do house prices respond to changes in these rules?

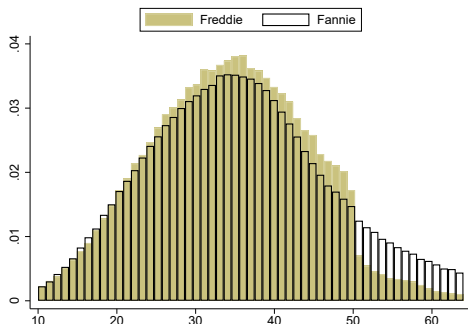
*Important for understanding the boom, effects of regulation.*



# Approach

- Lenders rely on software created by Fannie Mae or Freddie Mac.
- In 1999 Freddie tightened DTI rules (not public).

$$\text{DTI} = \frac{\text{mortgage payment} + \text{other obligations}}{\text{income}}$$



- Counties affected differently, depends on lender ties to Freddie.

# Main result

Debt-to-income limits have a large effect on house prices:

1. Consistent with response of constrained households in short run.
2. Continues to build over several years, suggesting feedback.

① Institutional Background

② Data and descriptive statistics

③ Policy change

④ Results

⑤ Model

## Institutional Background

# Automated underwriting software

- Software:

- Freddie's Loan Prospector (LP)
- Fannie's Desktop Underwriter (DU)

- Determines if Fannie or Freddie will buy a mortgage:

- Public rule: loan < conforming limit (\$453,100 in 2018).
- **Proprietary rules** relating to income, collateral and credit score.

- Could also be used for subprime/jumbo loans:

- “[Fannie and Freddie are] promoting the use of DU and LP for such non-conforming non-agency loan types as jumbos and subprime loans.” *Mortgage Banking, 1999*

# Software differences lead to local variation in DTI policy

Use Freddie Mac county market share from before the change.

Relationships exclusive and persistent.

*Mortgage Banking, 1999:*

- “It’s very **expensive to do both** [Fannie and Freddie’s software]. There’s the upfront costs and there’s all kinds of ancillary costs . . . So most lenders are opting to go with one based on where they have their primary business relationship.”
- “As soon as one comes out with something, it’s usually just a matter of time before the other does too. In the end they’re pretty close overall. I’m not sure every correspondent, broker or lender really needs both systems. There’s **tremendous overlap** and the product differentiation between the two is not a huge issue.”



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# Data and descriptive statistics

# Data

## 1. Document policy change:

- GSE Single Family Loan Performance
  - DTI, LTV, credit score for loans GSEs purchased.
  - Available from 1999.
- GSE Public Use Database
  - Loan-to-income, LTV for loans GSEs purchased.
  - Available from 1993.

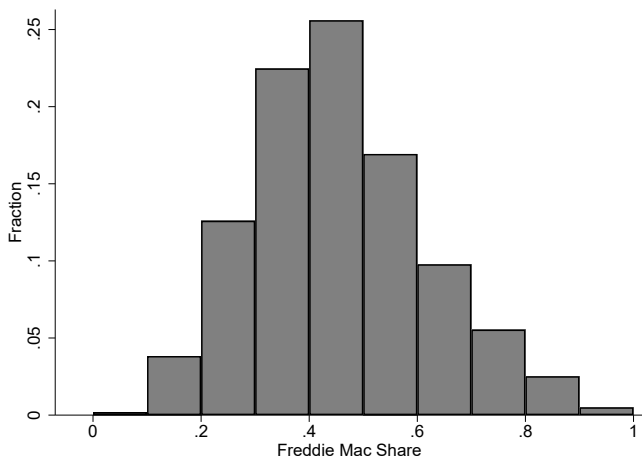
## 2. Calculate local exposure to Freddie Mac:

- HMDA
  - Loan-to-income, census tract, lender, was loan sold to Fannie/Freddie.
  - Available from 1991.

## 3. Measure effect on house prices

- CoreLogic county house price index.
- Similar results using FHFA house prices.

# County exposure to Freddie Mac (1998)



$$\text{Freddie}_c = \frac{\# \text{ Loans in county } c \text{ sold to Freddie}}{\# \text{ Loans in county } c \text{ sold to Freddie or Fannie}}$$

# Freddie exposure and other variables

- Counties are similar with respect to a number of variables (average DTI, underserved share, subprime share, share sold to Fannie/Freddie)
- Hower, high Freddie counties:
  - less coastal
  - less densely populated
  - lower median income

**Approach:** Within state variation, include controls and show divergence in prices coincides with policy [also similar results with reweighting].

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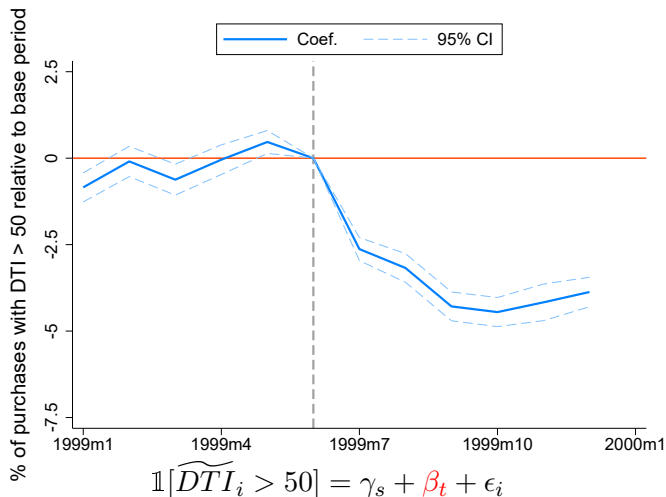
④ Results

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Policy change

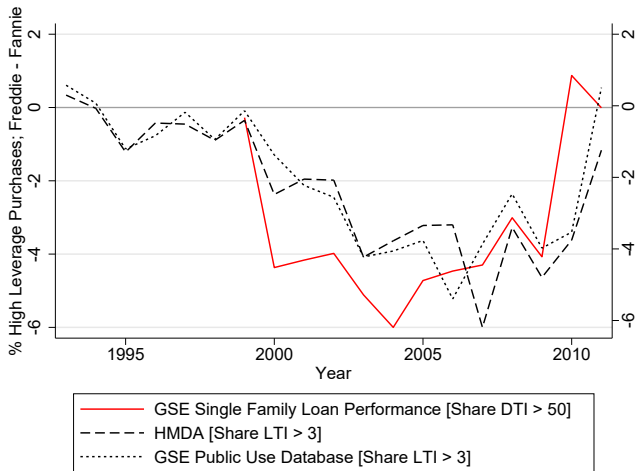
# Timing of change

Change is not publicly announced  $\rightarrow$  rely on data:

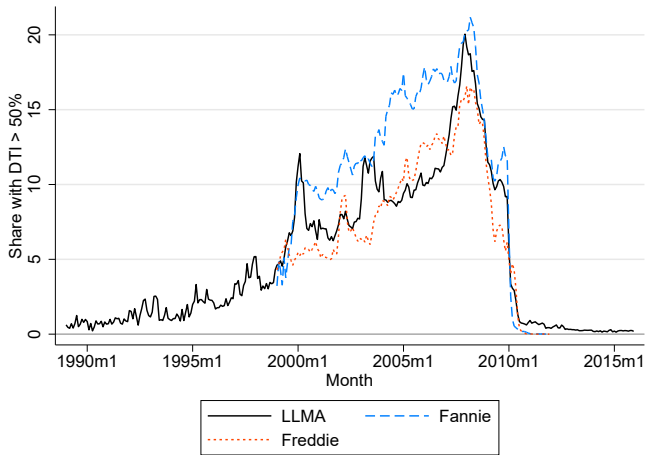




# Freddie applied tighter DTI rules until after the crisis



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# Results

# Outline

Show that in more exposed counties (relative to less exposed):

1. High DTI lending declines.
2. House prices decline.
3. Price decline continues for several years after change.

# High DTI share declines in more exposed counties

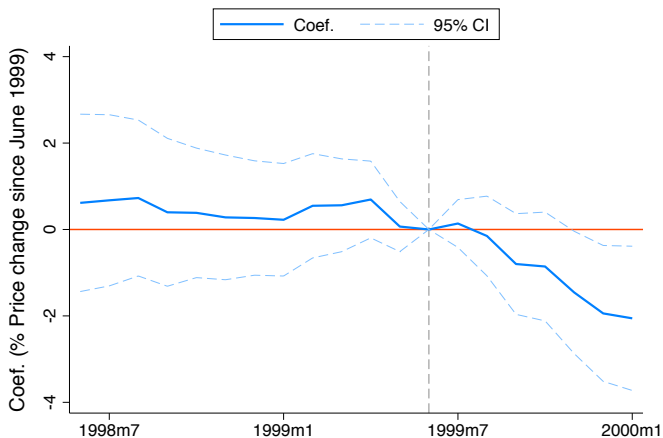
$$\text{High DTI}_{c,t} = \gamma_c + \gamma_{s,t} + \beta \text{Post}_t \cdot \text{Freddie}_{c,1998} + \alpha \text{Post}_t \cdot \text{Controls}_c + \epsilon_{c,t}$$

	Share DTI > 50	
	(1)	(2)
Post × Freddie	-3.79*** (1.21)	-2.80** (1.23)
County FE	X	X
State-Post FE	X	X
Controls		X
Number of Counties	1,197	1,195
Number of States	50	50
Number of Observations	2,394	2,390

Pre: Jan 1998 – Jun 1999; Post: Jul 1999 – Dec 2000.

# Debt-to-income tightening reduces house prices

$$\log(\text{Price}_{c,t}) = \gamma_c + \gamma_{s,t} + \beta_t \text{Freddie Share}_{c,1998} + \epsilon_{c,t}$$



# House prices decline in more exposed counties

$$\Delta \log(\text{Price}_c) = \gamma_s + \beta \text{Freddie share}_{c,1998} + \alpha \text{Controls}_c + \epsilon_c$$

	Jun 1999 – Dec 1999	
	(1)	(2)
Freddie Share	-2.48*** (0.78)	-1.94** (0.80)
State FE	X	X
Controls		X
Number of Counties	996	996
Number of States	49	49
Number of Observations	996	996



# Relative decline continues for several years

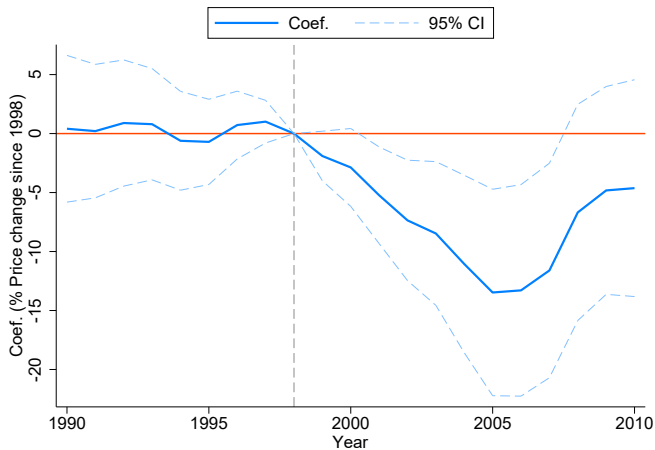
$$\Delta \log(\text{Price}_c) = \gamma_s + \beta \text{Freddie share}_{c,1998} + \alpha \text{Controls}_c + \epsilon_c$$

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	Jun 1999 – Dec 2002	
	(1)	(2)
Freddie Share	-8.93*** (2.72)	-7.79*** (2.57)
State FE	X	X
Controls		X
Number of Counties	996	996
Number of States	49	49
Number of Observations	996	996

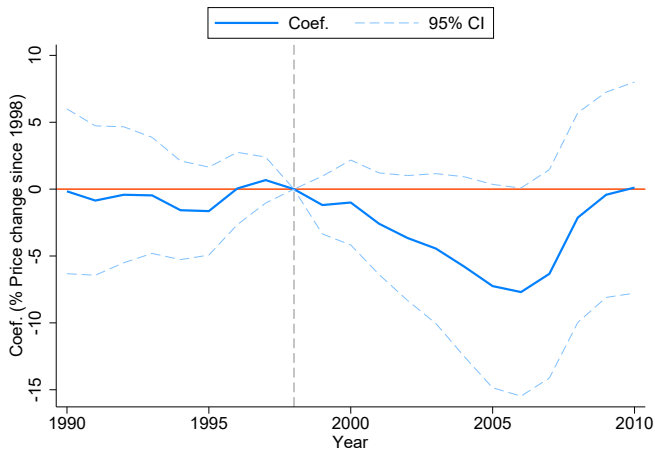
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# House price response: 1990 – 2010



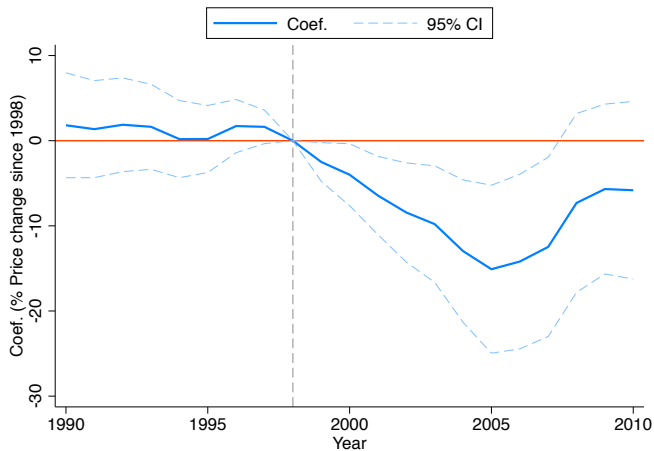
$$\log(\text{Price}_{c,t}) = \gamma_c + \gamma_{s,t} + \beta_t \text{Freddie Share}_{c,1998} + \alpha_t \text{Controls}_c + \epsilon_{c,t}$$

# House price response excl. top 20 CBSAs: 1990 – 2010



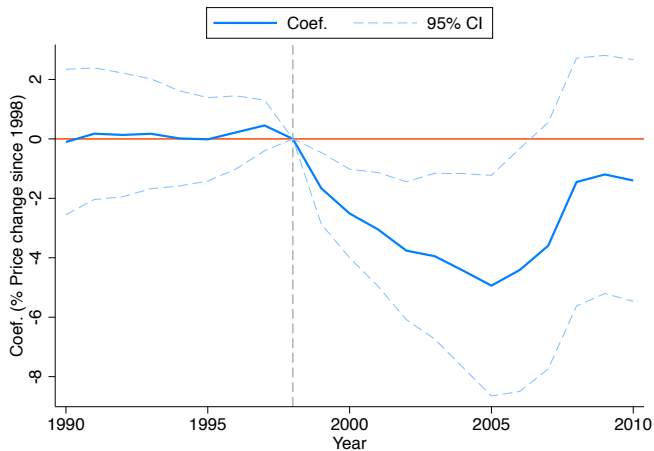
$$\log(\text{Price}_{c,t}) = \gamma_c + \gamma_{s,t} + \beta_t \text{Freddie Share}_{c,1998} + \alpha_t \text{Controls}_c + \epsilon_{c,t}$$

# House price response excl. sand states: 1990 – 2010



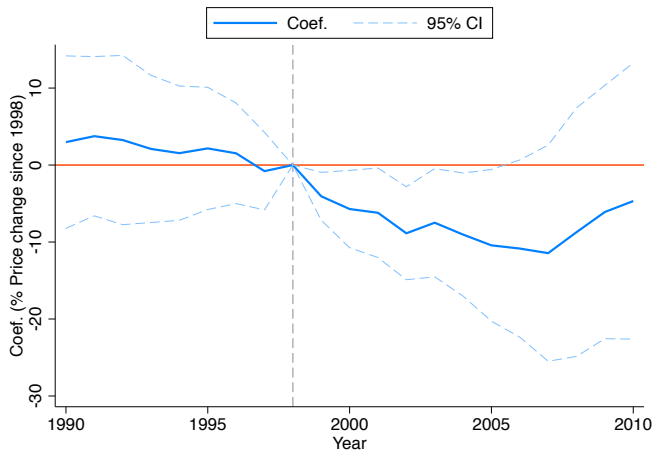
$$\log(\text{Price}_{c,t}) = \gamma_c + \gamma_{s,t} + \beta_t \text{Freddie Share}_{c,1998} + \alpha_t \text{Controls}_c + \epsilon_{c,t}$$

# House price response (binary; reweighted): 1990 – 2010



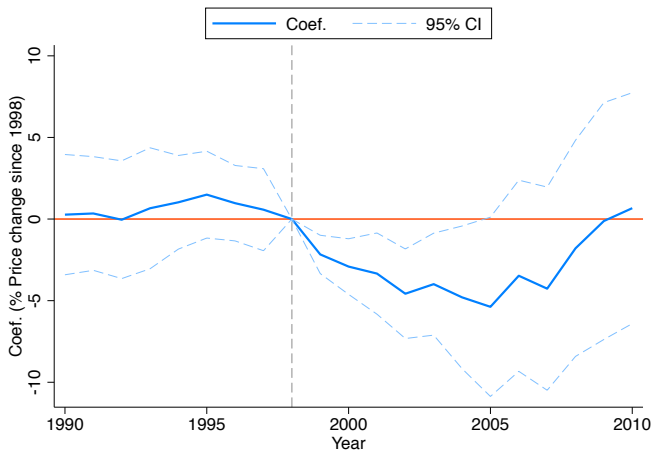
$$\log(\text{Price}_{c,t}) = \gamma_c + \gamma_{s,t} + \beta_t \text{High Freddie Share}_{c,1998} + \epsilon_{c,t}$$

# House price response (CBSA FE): 1990 – 2010



$$\log(\text{Price}_{c,t}) = \gamma_c + \gamma_{cbsa,t} + \beta_t \text{Freddie Share}_{c,1998} + \epsilon_{c,t}$$

# House price response (CBSA FE; binary; reweighted): 1990 – 2010



$$\log(\text{Price}_{c,t}) = \gamma_c + \gamma_{cbsa,t} + \beta_t \text{High Freddie Share}_{c,1998} + \epsilon_{c,t}$$

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Model

# Model of housing demand

- Receive utility from housing services and non-housing consumption:

$$u(H_i, C_i) = \alpha_i \log H_i + (1 - \alpha_i) \log C_i$$

- Allocate income across both, given cost of housing services:

$$y_i = C_i + uPH_i$$

*u* is user cost (*interest rate + property tax + depreciation - price growth*)

**But:** must buy the housing asset (at price  $P$ ) to consume housing.

Means choice is restricted by available downpayment and income in presence of LTV and DTI constraints.

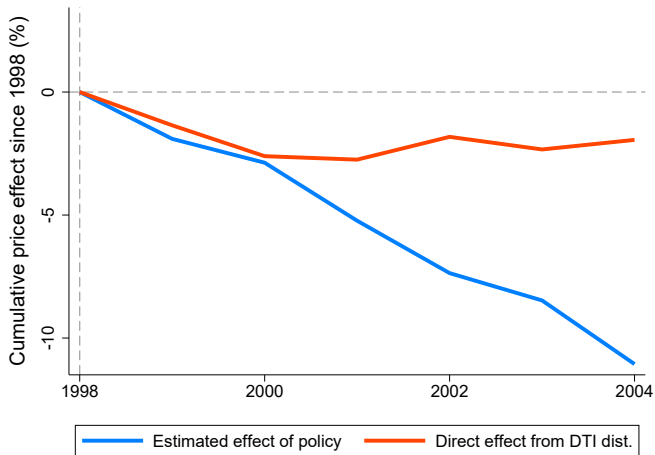
# Back-of-the-envelope formula

If households above the new DTI cutoff (50%) respond by cutting loan size (i.e. holding income and downpayment fixed):

$$\% \Delta P \approx \frac{P(\text{constrained by DTI})}{1 + \epsilon} \frac{\bar{y}}{f(r)P\bar{H}} \left( 0.5 - \frac{\sum_i DTI_i \mathbb{1}[DTI_i > 50]}{\sum_i \mathbb{1}[DTI_i > 50]} \right)$$

- $f(r)$  is fixed rate payment per \$1 of debt,  $r$  is Freddie 30-year rate.
- Compute constrained share using diff. between Fannie and Freddie share  $DTI > 50$ , times share using mortgage.
- Use median income for  $\bar{y}$  and median house price for  $P\bar{H}$ .
- $\epsilon$  is the housing supply elasticity.
- Compute mean DTI conditional on  $DTI > 50$  using Fannie data.

# Comparison with empirical results



# Effect on unconstrained households

More exposed locations have weaker house price history.

With adaptive expectations, user cost is higher ( $u = r + \tau + \delta - g$ ):

- Calibrate a rule to match Case, Shiller & Thompson (2012):  
 $g = A(\lambda) \sum_{j=0}^{t-t_0} (1 - \lambda)^j g_{t-j}$  where  $\lambda = 0.11$ .
- Compute  $g$  adjusting for policy effect and get  $\% \Delta$  user cost (difference between exposed and unexposed areas).

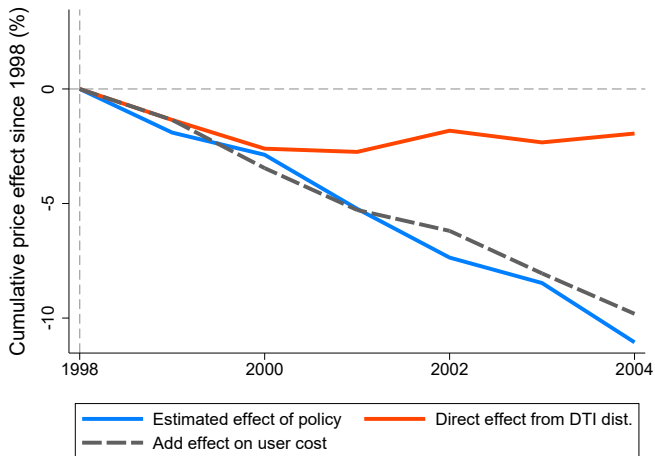
# Effect on unconstrained households

More exposed locations have weaker house price history. With adaptive expectations, user cost is higher ( $u = r + \tau + \delta - g$ ):

$$\% \Delta P \approx \frac{1}{1 + \epsilon} \left[ -P(\text{responds to user cost change}) \cdot \% \Delta \text{user cost} \right. \\ \left. + P(\text{constrained by DTI}) \cdot \frac{\bar{y}}{f(r)P\bar{H}} \left( 0.5 - \frac{\sum_i DTI_i \mathbb{1}[DTI_i > 50]}{\sum_i \mathbb{1}[DTI_i > 50]} \right) \right]$$

- Use  $\delta = 2\%$ ,  $\tau = 1.2\%$ ,  $\epsilon = 0.1$ .
- Assume high LTV and affected groups with  $DTI > 50$  do not respond to user cost change.

# Comparison with empirical results



# Conclusion

- Debt-to-income policies have a large effect on house prices.
- The effect builds over time.
- Not just policy: policy + expectations?